

1 **Differentiation of Self and Interpersonal Functioning with the Level of Personality**
2 **Functioning Scale – Brief Form 2.0 (LPFS-BF 2.0)**

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29

30 **Abstract**

31 Research on Criterion A of the Alternative Model for Personality Disorders is recently
32 expanding and provides mixed results concerning the unidimensional operational
33 definition of severity by the model, characterized by impaired self (identity and self-
34 direction) and interpersonal (empathy and intimacy) functioning. Studies resulted in
35 one, as well as two or more factor structures. The present study demonstrated the
36 importance of the structural and relational differentiation of self and interpersonal
37 dimensions of personality functioning. 1074 participants (community and clinical mixed
38 sample) completed the Level of Personality Functioning Scale – Brief Form 2.0 (LPFS-
39 BF 2.0), the Personality Inventory for DSM-5 Short Form and the Questionnaire for the
40 World Health Organization Disability Assessment. An LPFS-BF 2.0 two-factor
41 structure with self and interpersonal functioning factors was corroborated by
42 confirmatory factor analyses and bifactor modeling. Joint Exploratory Factor Analysis
43 of the LPFS-BF 2.0 domains with maladaptive personality domains clearly
44 differentiated the personality functioning factors. While the self functioning factor was
45 more closely linked to negative affect (and to disinhibition and psychoticism), the
46 interpersonal functioning factor connected to detachment. Self functioning predicted
47 functional impairment along and beyond personality domains. The LPFS-BF 2.0
48 appears a useful tool for clinical routine monitoring of both self and interpersonal
49 functioning.

50

51

Background

52 Section III of the fifth edition of the Diagnostic and Statistical Manual of Mental
53 Disorders (DSM-5; (2013)) introduced the Alternative Model for Personality Disorders
54 (AMPD) to respond to criticism on categorical approaches in the field of personality
55 pathology (Widiger & Trull, 2007). AMPD is based on a two-step dimensional
56 assessment of personality disorders (PDs): level of personality functioning (LPF)
57 (Criterion A) and maladaptive traits (Criterion B) (APA, 2013). Criterion A assesses the
58 overall severity of personality functioning, characterized by central components of
59 personality impairment, self (identity and self-direction) and interpersonal (empathy and
60 intimacy) functioning. Criterion B defines specific individual differences in the trait
61 profile. Since the publication of DSM-5, much of the AMPD literature has focused on
62 providing validity evidence for the 5 trait domains (negative affect, detachment,
63 antagonism, disinhibition, and psychoticism) and its 25 facets defined in Criterion B
64 (Al-Dajani et al., 2016; Somma et al., 2019; Watters & Bagby, 2018). More recently,
65 scientific attention has shifted towards Criterion A, with a considerable set of mixed
66 results both supporting and disagreeing with the unidimensional operational definition
67 of severity by this model, as well as with its specific contribution to the diagnosis of
68 PDs and prediction of related outcomes (Sharp & Wall, 2021; Sleep et al., 2021;
69 Widiger et al., 2019; Zimmermann et al., 2019).

70 In order to assess DSM-5 Criterion A, the Level of Personality Functioning
71 Scale (LPFS) was launched (APA, 2013; Bender et al., 2011). LPFS defines 12
72 elements by assessing three facets within identity, self-direction, intimacy, and
73 empathy. Five distinct levels of impairment are defined in each element, from no
74 impairment (Level 0) to extreme impairment (Level 4). This results in 60 descriptions
75 of personality dysfunction (12 elements X 5 levels) in order to determine PD severity on

76 a general dimension. Level 2 (moderate severity) is the threshold for presence of a PD.
77 Additionally to the original clinician-rated LPFS (Bender et al., 2011), other structured
78 interviews (Bender, 2018; Hutsebaut et al., 2017) and self-report measures designed to
79 assess LPF have been developed (Huprich et al., 2018; Morey, 2017; Siefert et al.,
80 2020). One of those instruments is the Level of Personality Functioning Scale-Brief-
81 Form (LPFS-BF) (Hutsebaut et al., 2016), a quick screener of personality functioning
82 with 12 dichotomous (yes/no) items with each LPFS element represented by one item.
83 Its most recent revision, the LPFS-BF 2.0 (Weekers et al., 2019) converted the
84 dichotomous items into a 4-point Likert scale, to be more aligned with the dimensional
85 approach of AMPD and to facilitate sensitivity for measuring change in personality
86 functioning during treatment.

87 The unidimensional structure of personality functioning, capturing the core of
88 personality pathology, as originally intended (Bender et al., 2011; Morey, 2019; Sharp
89 & Wall, 2021) has been evidenced (Cruitt et al., 2019; Gamache et al., 2019; Hopwood
90 et al., 2018; Morey, 2017). Although the four subdimensions defined in criterion A
91 (identity, self-direction, empathy and intimacy) also have some support (McCabe et al.,
92 2021; Sleep et al., 2019), and even a six-factor solution when personality functioning is
93 evaluated in a disorder-specific manner (Anderson & Sellbom, 2016) has been reported.
94 The instrument used to measure personality functioning may have influenced on the
95 different results in terms of structure. Yet, empirical evidence most often endorsed the
96 two-factor model (with two self and interpersonal factors strongly correlated) on
97 different self- and other-report measures and expert ratings (Bliton et al., 2022;
98 Hutsebaut et al., 2016; Roche et al., 2018; Spitzer et al., 2021; Stover et al., 2020;
99 Weekers et al., 2019; Zimmermann et al., 2015). Studies in patients using structured
100 interview measures also confirmed two factors (Heissler et al., 2021; Hummelen et al.,

101 2021; Ohse et al., 2022). The current study will focus on the LPFS-BF 2.0. The original
102 validation study by Weekers et al. (2019) in patients required some modifications to
103 reach adequate fit for a two-factor structure (i.e. a correlated error term and a cross-
104 loading of item 11). However, confirmatory factor studies with LPFS-BF 2.0 in non-
105 clinical samples also gathered more support for a two-factor (Lakuta, 2022; Le Corff et
106 al., accepted 2022; Natoli et al., accepted, 2022) than for an unidimensional model
107 (Weekers et al., 2022). Yet, further evidence on the LPFS-BF 2.0 structure, especially in
108 clinical samples, is needed.

109 Although self and interpersonal factors are often strongly intercorrelated, we
110 consider their differentiation of substantive interest, given the distinct association with
111 personality traits. Correlational studies have reported greater association of self
112 functioning subdomains with neuroticism and conscientiousness adaptive personality
113 domains, while interpersonal functioning correlated in a greater extent with
114 agreeableness (Berghuis et al., 2014). Similarly, self functioning (and its subdomains)
115 related more strongly with maladaptive negative affect domains (with also associations
116 with detachment and psychoticism) while interpersonal functioning was linked to
117 antagonism and detachment (Few et al., 2013; Lakuta, 2022). Studies using exploratory
118 structural equation (ESEM) models have also shown that the self impairment dimension
119 of identity grouped with neuroticism and negative affect domains (McCabe et al., 2021;
120 Oltmanns & Widiger, 2016). Clinician-rated based traits of depressivity and separation
121 insecurity appear on the same factor than self functioning, while interpersonal
122 functioning appeared to group with antagonistic traits such as grandiosity and
123 callousness and with detachment traits such as intimacy avoidance or withdrawal
124 (Lakuta, 2022; Zimmermann et al., 2015). These empirical results indicate the
125 importance to further examine the distinctiveness of self and interpersonal subdomains

126 of personality functioning in their relation with personality domains (Roche & Jaweed,
127 2021).

128 Another important debate concerns the incremental predictive ability of
129 Criterion A above Criterion B of AMPD and vice versa to explain or predict related
130 outcome measures. Given the empirical overlap between both constructs (measures of
131 Criterion A and B have been found to be highly correlated), some question if this
132 conceptual differentiation can hold (Zimmermann et al., 2019). Other argue they are
133 still distinct given their incremental validity for important outcomes. Most attention has
134 focused on the incremental validity for predicting PD categories (Anderson & Sellbom,
135 2018; Few et al., 2015; Nysaeter et al., 2022; Sleep et al., 2019). A recent study (Hobbs
136 et al., 2023) examining the relationship of personality traits with quality of life and
137 impairment, indicated that it is the pathological pole of traits that is associated with
138 lower quality of life and increased impairment. Also, there is strong evidence that
139 personality pathology has a strong impact on basic social, occupation and global
140 functioning, even in a greater extent than other psychopathology (Skodol, 2018),
141 nevertheless scarce studies are available studying how each dimension of Criterion A
142 contributes to the explanation of psychosocial functioning and disability. Two studies
143 (including a clinical and a mixed clinical-community sample) suggested an increment
144 for LPFS ratings in predicting psychosocial functioning compared with total number of
145 DSM-IV criteria (Morey et al., 2013), with specifically a greater contribution of the self
146 functioning factor in the study by Buer Christensen et al. (2020). Congruently, in a
147 borderline PD study, identity (a subfactor of self functioning) has been found to be the
148 strongest predictor for psychosocial impairment functioning (Esguevillas et al., 2018).
149 Also, Weekers and colleagues (2019) found the self functioning domain to be more
150 sensitive to change after 3 months of inpatient treatment than the interpersonal domain.

151 In non-clinical samples quality of life was also taken into consideration.
152 Weekers and colleagues (2022) reported significant associations between LFPPF-BF 2.0
153 global score and quality of life and social and occupational impairment. Similarly,
154 Lakuta (2022) also pointed towards self functioning as a significant predictor of well-
155 being, above and beyond maladaptive personality traits. Further evidence determining
156 which elements of AMPD are most related with impairment in clinical samples can
157 contribute to identify predictors and moderators of treatment outcomes (Zimmermann et
158 al., 2019). In addition, as noted above, the studies conducted to date mostly involve
159 specific patient or community samples. In this sense, the inclusion of mixed samples
160 would allow a more adequate representation of the entire continuum under evaluation.

161 Considering the above mentioned, and since self and interpersonal functioning
162 were recently defined as key general predictors of PD based on a summary of empirical
163 studies (Zimmermann et al., 2022), the general objective of the present study is to
164 provide additional evidence on the differentiation of self and interpersonal dimensions
165 of personality functioning. More specifically using a quick screener, the LPFS-BF 2.0,
166 in a mixed community and clinical sample we aim to: 1) provide structural evidence for
167 a two factor model with self and interpersonal as related dimensions; 2) obtain evidence
168 of differential associations of self and interpersonal functioning with maladaptive
169 personality domains; 3) determine the contribution self and interpersonal functioning
170 for explaining functional impairment, and their incremental validity above personality
171 domains.

172 We expect to find a two-factor structure with a high correlation between the self
173 and interpersonal factor (Lakuta, 2022; Le Corff et al., accepted 2022; Natoli et al.,
174 accepted, 2022). When jointly factor analyzing personality functioning with
175 maladaptive personality domains, we hypothesize self functioning to be grouped with

176 negative affect (and possible also connect to detachment and psychoticism), while
177 interpersonal functioning to be more related to the detachment and antagonism domains
178 (Few et al., 2013; Sleep et al., 2019; Sleep et al., 2020; Zimmermann et al., 2015). We
179 assume that degree of personality dysfunctioning will be related to level of functional
180 impairment, given the earlier found associations with social and occupational
181 impairment (Weekers et al., 2022). Finally, we expect self functioning to be more
182 predictive for functional impairment than interpersonal functioning, given it was more
183 strongly related to disability and psychosocial impairment (Buer Christensen et al.,
184 2020; Esguevillas et al., 2018; Zimmermann et al., 2019).

185

186

Materials and methods

187 Participants and procedure

188 In order to well capture the personality continuum by a wide variability in the
189 manifestation of maladaptive personality traits and personality functioning, the present
190 study was conducted in a mixed sample (N = 1074) composed of community adults (n =
191 717) and patients (n = 357). Inclusion criteria for the community sample were being
192 between 18 and 80 years old and not having a mental disorder. The study was approved
193 by the Bioethics Committee of the Province of Huelva (Junta de Andalucía, Spain) (No.
194 PI 040/18). All participants gave informed consent to anonymously and voluntary
195 participate. Inclusion criteria for the clinical sample were being adult (≥ 18) and being
196 under treatment in a mental health service during data collection. Exclusion criteria in
197 both samples were: 1) having been diagnosed with a medical or psychological disorder,
198 such as active psychosis or severe mental retardation, that disqualifies from taking the
199 tests, or 2) not signing the informed consent form.

200 An online access panel provider recruited a sampling frame of community
201 participants, comprising 155,000 adults representative of the overall Spanish
202 population. This company, specialized in online data collection, is certified with the
203 International Organization for Standardization (ISO) 26362 quality standards. For
204 present research, they provided the community sample group of the current sample
205 selecting them by a stratified random sampling approach with proportional allocation of
206 panelists according to gender, age (range: 18 - 75 years), and geographical region of the
207 Spanish territory. Questionnaires were administered online. Only participants who fully
208 completed all measures were retained. Less than 3% of the invited participants provided
209 incomplete answers. The sample composition of the current study thus mirrored the
210 sociodemographical distribution of the Spanish population. Among community
211 participants, 44.4% were women, with ages ranging from 18 to 75 years ($M = 46.30$; SD
212 $= 14.65$). 61.6% were employed, and 3.3% had primary education, 49.8% had
213 secondary education, and 46.9% had completed full bachelor.

214 The clinical sample was collected via clinicians from mental health services in
215 the Spanish province of Huelva by using of a non-probabilistic procedure. Patients were
216 diagnosed by clinicians according to DSM diagnostic criteria (at time of data collection,
217 DSM-IV was used in clinical practice). These diagnoses were only used to determine
218 whether patients were eligible for the clinical group, the specific diagnoses were not
219 used in the analyses in present research. Clinical patients completed the tests in rooms
220 set up in the centers where they were recruited. The most frequent diagnostic categories
221 among patients were Depressive Disorders (35.87%), Anxiety Disorders (31.31%) and
222 Trauma and Stressor-Related Disorders (23.40%). 9.5% of patients fulfilled criteria for
223 a Personality Disorder. Participants in this sample consisted of 63.0% women and
224 ranged in age from 18 to 80 years ($M = 37.72$; $SD = 14.48$); 35.0% were employed at

225 the time of the study, and 13.7% had primary education, 68.1% had secondary
226 education, and 18.2% had completed higher education.

227

228 **Instruments**

229 *Level of Personality Functioning Scale - Brief Form 2.0* (LPFS-BF 2.0; Weekers et al.,
230 2019). The Spanish version of LPFS-BF 2.0 (available for Spain on the website of the
231 original authors: [LPFS-BF Scale - ggz de Viersprong](#)), was used. This version showed
232 adequate reliability and criterion validity (Le Corff et al., accepted 2022). The LPFS-BF
233 2.0 consists of two domains: Self functioning (6 items measuring subdomains of
234 identity and self-direction) and Interpersonal functioning (6 items evaluating
235 subdomains of empathy and intimacy), measured on a 4-point Likert scale ranging from
236 1 (completely untrue) to 4 (completely true). The LPFS-BF-2.0 total score ranges from
237 12 to 48 and the two domain scores can range from 6 to 24. The higher the score, the
238 greater the deficits in self and interpersonal functioning. In the present sample, internal
239 consistency values of total scale and domains were adequate: Total scale ($\alpha = .87$; $\Omega =$
240 $.88$; MIC = $.34$); Self functioning ($\alpha = .88$; $\Omega = .89$; MIC = $.54$); Interpersonal
241 functioning ($\alpha = .69$; $\Omega = .70$; MIC = $.28$)

242

243 *Personality Inventory for DSM-5 Short Form* (PID-5-SF; Maples et al., 2015). The
244 Spanish version translated and validated by Diaz-Batanero et al. (2019) was used. The
245 items of this instrument are grouped into the 25 facets as described in section III of the
246 DSM-5 and the five major scales correspond to the DSM-5 trait domains: Negative
247 affect, Detachment, Antagonism, Disinhibition, and Psychoticism. This inventory
248 assesses pathological personality traits in adults. It is composed of 100 items with a
249 Likert-type format with four options: strongly disagree, moderately disagree,

250 moderately agree, and strongly agree. In the present study, Cronbach's alpha internal
251 consistency coefficients ranged from .86 (Negative affect) to .88 (Detachment and
252 Antagonism) ($\Omega = .83-.88$; MIC = .30-.48).

253

254 *Questionnaire for the World Health Organization Disability Assessment (WHODAS*
255 *2.0)* (Ustün et al., 2010). The 12-item Spanish brief version was used, measuring
256 functional impairment in six domains of life: cognition, mobility, self-care, getting
257 along with other people, life activities and participation in society. This short version
258 provides a single global index of impairment in the past 30 days (e.g. “*In the past days,*
259 *how much difficulty did you have in taking care of your household responsibilities*”, “*In*
260 *the past days, how much difficulty did you have in dealing with people you do not*
261 *know*”). In the present study, internal consistency coefficients for the composite scores
262 were $\alpha = .89$; $\Omega = .88$; MIC = .45.

263

264 *Sociodemographic variables.* Questions on sex, age, educational level, current
265 employment status, marital status, etc. were included.

266

267 **Data analysis**

268 Less than 0.1% of item data was missing. Descriptive statistics (M and SD) for scores
269 on all scales were computed. Reliability coefficients (McDonald's omega, Cronbach's
270 alpha and mean inter-item correlations (MIC)) were computed. T-tests were performed
271 to assess mean differences across the clinical and non-clinical samples. Effect sizes
272 were quantified using Cohen's *d*. Values greater than |0.20|, |0.50|, and |0.80| represent
273 respectively small, medium, and large effect sizes, (Cohen, 1992).

274 The structure of LPFS 2.0 was estimated with Confirmatory Factor Analyses.
275 Two models were tested: *Model 1*, with all items loading on a general factor of
276 personality functioning; *Model 2*, with two correlated factors representing self and
277 interpersonal functioning. The CFA analyses were conducted using the WLSMV robust
278 estimation method which does not assume normally distributed variables and provides
279 the best option for modelling categorical or ordered data (Brown, 2015). To evaluate
280 goodness of fit for both models, the following fit statistics were used: Comparative Fit
281 Index (CFI), Tucker-Lewis Index (TLI), Root Mean Square Error Approximation
282 (RMSEA), and Standardized Root Mean Square Residual (SRMR). Values of CFI and
283 TLI above .95, and values of RMSEA and SRMR below .08 and .05, respectively, were
284 considered indicative of good fit (Hu & Bentler, 1999).

285 To examine the dimensionality of LPFS-BF 2.0, bifactor modeling was applied.
286 The Bifactor model (Model 3) included a general factor explaining covariance across all
287 items and two orthogonal specific factors (self functioning and interpersonal
288 functioning) explaining excess shared variance among item clusters. Despite the critique
289 that bifactor might result in biased fit estimates (Eid et al., 2017; Heinrich et al., 2021),
290 they are recommended to evaluate dimensionality and in this way complement
291 correlated factors analysis (Rodriguez et al., 2016a). Information of the bifactor model
292 was used to calculate omega, omega hierarchical (omegaH), and explained common
293 variance (ECV) (Rodriguez et al., 2016b). ECV characterizes the proportion of common
294 variance across items that is explained by each factor. For the general factor, a value
295 greater than .70 indicates that the majority of common variance is explained by the
296 general factor (Reise et al., 2013). Omega reflects the proportion of total score variance
297 that can be attributed to common factors. For the general factor, all items are
298 considered; for specific factors, only items loading on that factor are included. OmegaH

299 is the proportion of the total score variance that can be attributed to the general factor
300 after accounting for all other specific factors. An $\omega_H > .80$ indicates that the total
301 score can be interpreted as unidimensional (Reise et al., 2013; Rodriguez et al., 2016a).

302 Next, the joint structure of LPFS-BF 2.0 and PID-5-SF was explored. Parallel
303 analysis (PA) for principal axis factoring using permutations of the raw data (Horn,
304 1965; O'Connor, 2000) was performed to determine the number of factors to retain. PA
305 tends to indicate more factors, especially when factors are highly correlated (Buja &
306 Eyuboglu, 1992), therefore Auerswald and Moshagen (2019) recommend using multiple
307 criteria. Empirical Kaiser Criterion (Braeken & van Assen, 2017), Comparison Data
308 (CD) (Ruscio & Roche, 2012) and fit statistic differences (Finch, 2020) were estimated
309 as additional procedure to determine the number of factors to retain. An EFA with three
310 factors was estimated using a maximum likelihood extraction and oblimin rotation.
311 Factor loadings $\geq .30$ were considered meaningful. A factor loading is basically the
312 correlation coefficient for the variable and factor, and according to Cohen (1992) .30 is
313 a medium effect size.

314 In order to determine if functional impairment differs according to level of
315 personality functioning, cutoffs provided by Weekers et al. (2022) for the LPFS-BF 2.0
316 total score were used to classify participants into impairment levels: no indication of
317 impairment (total score LPFS-BF 2.0 < 26), subclinical impairment ($26 \leq$ total score
318 LPFS-BF 2.0 ≤ 36), and clinical impairment (total score LPFS-BF 2.0 >36). Groups
319 generated this way were compared on functional impairment measured by WHODAS
320 2.0 with an ANOVA.

321 Then, to evaluate the predictive validity of each LPFS-BF 2.0 scale in explaining
322 functional impairment as measured by the WHODAS 2.0., and their incremental
323 validity above PID-5 scales, a structural equation modelling (SEM) approach was

324 considered, as suggested by Wang & Eastwick (2020). Prior to conducting the SEM, a
325 power analysis was performed. Using an expected effect size of the increment of LPFS-
326 BF 2.0 beyond PID-5 scales of .05 (Lakuta et al., 2022), an alpha level of .05 and a
327 sample size of 1074, the statistical power showed a value of .26 for Self functioning and
328 .27 for Interpersonal functioning (Wang & Rhemtulla, 2021) (see supplement, table S1).

329 These values did not reach the suggested power of e.g. 80% (or more), so with
330 the available sample size the analysis cannot be done using a SEM approach (Wang &
331 Eastwick, 2020). Therefore, multiple linear regressions, controlling for age and sex,
332 were employed. A Bonferroni correction for multiple testing was applied by only
333 interpreting results at p-level of $p < .0027$ (result of dividing alpha level .05 by the
334 maximum number of parameters estimated in the regression model).

335 CFA was completed in Mplus (Version 8; Muthén & Muthén, 2019), joint EFA
336 was computed in JASP and number of factors to be extracted were decided using
337 EFA.dimensions and EFA.tools R (Steiner & Grieder, 2020) packages. Other analyses
338 were performed with SPSS 27 (IBM, 2021).

339

340

Results

341 *Descriptive statistics*

342 Table 1 shows the descriptive statistics (M and SD) of the scores on all measured scales
343 in each subsample and the total sample.

344

*** INSERT TABLE 1 HERE ****

345 The clinical sample obtained significantly higher LPFS-BF 2.0 total, Self functioning
346 and Interpersonal functioning scores, with effect sizes being moderate to high ($d = 0.44$
347 to 1.08). For the other measures (PID-SF and WHODAS 2.0), t-tests showed mean
348 differences across the clinical and non-clinical samples for all scales except for PID-5

349 attention Seeking, Callousness, Deceitfulness, Grandiosity, Manipulativeness, and
350 Antagonism domain. Cohen's d ranged from 0.03 (PID-Callousness) to 1.17
351 (WHODAS 2.0) (with large effect sizes for most scales). These differences observed
352 between the two samples show that the general mixed sample indeed allows a broader
353 representation of the whole continuum of personality functioning and therefore
354 subsequent results are presented for the mixed total sample.

355

356 ***Confirmatory factor analyses of the LPFS-BF 2.0***

357 The two correlated-factor model provided best fit (Table 2).

358 *** INSERT TABLE 2 HERE ****

359

360 Dimensionality parameters (calculated with the bi-factor model) supported a
361 multidimensional structure (ECV = .51; $\omega = .84$, $\omega_H = .59$) for the LPFS-BF
362 2.0, with specific factors accounting for most of the common variance. Both specific
363 factors demonstrated reliability ($\omega_{\text{self}} = .83$; $\omega_{\text{interpersonal}} = .69$) and
364 captured reliable variance beyond the general factor ($\omega_{H\text{self}} = .35$;
365 $\omega_{H\text{interpersonal}} = .31$). Self and Interpersonal functioning factors were largely and
366 significantly associated ($r = .72$) in the two correlated-factor models (see Figure 1).

367 *** INSERT FIGURE 1 HERE ****

368

369 ***Exploratory joint structure of the LPFS-BF 2.0 (Criterion A) and PID-5-SF***

370 ***(Criterion B) scales***

371 PA indicated a maximum of three factors to be retained (see supplement, table S2).

372 EKC (see supplement, table S3) and CD agreed retaining 3 factors (see supplement,

373 figure S1). Fit indices for different factor solutions of the exploratory factor analysis are
374 displayed in table 3.

375 *** INSERT TABLE 3 HERE ****

376 Fit indices of the 3 factor solution showed adequate values (CFI = .998; TLI = .994;
377 SRMR = .005; RMSEA = .033; 90% RMSEA CI = .001 - .0068) and difference of
378 RMSEA between two and three factors suggest retaining three factors (Finch, 2020).
379 Table 4 shows standardized factor loadings of the LPFS factor and PID-5 domains
380 scales after oblimin rotation. F1 captured LPFS Self functioning along with PID-5
381 negative Affect, Disinhibition and Psychoticism; F2 accounted for the LPFS
382 Interpersonal functioning and PID-5 Detachment; F3 had Antagonism as main loading.
383 Correlation between factors ranged from .37 (F2-F3) to .69 (F1-F2).

384 *** INSERT TABLE 4 HERE ****

385

386 *Predictive and incremental validity of LPFS-BF 2.0 (Criterion A) and PID-5-SF*
387 *(Criterion B) scales to explain functional impairment as measured with the*
388 *WHODAS 2.0*

389 Considering cutoffs provided by Weekers et al. (2022), 928 participants (86.5%) had no
390 indication of impairment, 88 participants (8.1%) had an indication of subclinical
391 impairment, and 58 participants (5.4%) had a clinical impairment. Between the three
392 groups significant differences ($F_{(2,1069)} = 152.14, p < .001$) were observed on WHODAS
393 scores, with higher scores for the clinical group ($M = 37.43, DT = 9.58$) followed by the
394 subclinical group ($M = 28.59, DT = 10.24$) and the no impairment group ($M = 18.14,$
395 $DT = 7.69$).

396 Controlling for age and sex in step 1, results show that both LPFS-BF 2.0 scales,
397 significantly contributed to explain functional impairment measured by WHODAS 2.0

398 in step 2, ($F_{(4,1061)} = 178.84; p < .001$), reaching a $R^2 = .40$ (Model A) (see Table 5).
399 When explaining WHODAS 2.0 scores using PID-5 domains in the second step (Model
400 B), the model was also statistically significant ($F_{(7,1058)} = 124.33; p < .001$), with all
401 PID-5 domains scores contributing to the model, reaching a $R^2 = .45$, (see Table 6). In
402 both models, LPFS-BF 2.0 and PID-5 thus included together in a third step. All
403 variables remained significant, except for LPFS Interpersonal Functioning $F_{(9, 1056)} =$
404 $105.01; p < .001$). The addition of PID-5 domains beyond LPFS lead to an increment of
405 R^2 of 7% in model A. Adding LPFS in a third step incremented the explained variance
406 with an additional 2% in model B. Standardized coefficients with moderate effect size
407 (Peterson & Brown, 2005) corresponded to LPFS Self functioning ($\beta = .26; r = .31$) and
408 Detachment ($\beta = .24; r = .29$).

409 *** INSERT TABLE 5 HERE ****

410 *** INSERT TABLE 6 HERE ****

411

412 Discussion

413 The present work provides empirical evidence to differentiate two general
414 dimensions of criterion A, when operationalized through the LPFS-BF 2.0, and
415 demonstrates the differential contribution of self and interpersonal functioning in the
416 explanation of functional impairment. Although the two factors were highly
417 intercorrelated, results showed they can be differentiated from each other. When jointly
418 factored with the maladaptive personality domains, an intermingled structure between
419 functioning and maladaptive traits was found. Yet the results also displayed clear
420 differences between the personality functioning factors, with the self functioning factor
421 more closely linked to negative affect, beside connections with disinhibition and
422 psychoticism, and the interpersonal functioning factor linked to detachment. Moreover,

423 our results suggested a relevant role of self functioning in the explanation of the
424 functional impairment along and beyond personality facets. The main results and their
425 implications are further discussed below.

426

427 *Confirmatory factor analyses of the LPFS-BF 2.0*

428 The dimensionality analysis indicated that the specific factors of self and interpersonal
429 functioning have a higher contribution of variance compared to the general factor. Also,
430 the confirmatory factor results and fit indices support a two-dimensional
431 conceptualization of personality functioning for the LPFS-BF 2.0. These findings are
432 analogous to Bliton et al. (2022) in a sample of students. At the same time, a moderate-
433 high correlation was observed between the factors, similar, to other LPFS-BF 2.0
434 studies that have found correlations between self and interpersonal functioning ranging
435 from .46 to .69 (Bach & Hutsebaut, 2018; Lakuta, 2022; Weekers et al., 2019). This
436 implies individuals may differ to the extent their problems in personality functioning
437 maybe expressed more by self or interpersonal pathology (Zimmermann et al., 2022).
438 Present work provides further important evidence on the robustness of the LPFS-BF 2.0
439 empirical two-dimensional structure. Previous work using different criterion A
440 measures showed diverse evidences regarding the uni- versus multidimensional nature
441 and it was therefore recommended by Roche and Jaweed (2021) to expand the research
442 among the different LPFS instruments available. This study clearly corroborates the
443 greater consensus towards a the two factor structure for the instrument under study, like
444 observed with the original LPFS-BF (Bliton et al., 2022; Hutsebaut et al., 2016) and
445 current LPFS-BF 2.0 (Bach & Hutsebaut, 2018; Lakuta, 2022; Le Corff et al., accepted
446 2022; Weekers et al., 2019).

447

448 *Exploratory joint structure of the LPFS-BF 2.0 (Criterion A) and PID-5-SF*

449 *(Criterion B) scales*

450 Regarding the joint exploratory factor analysis of the LPFS-BF 2.0 and PID-5-
451 SF scales, present results are mostly congruent with our hypotheses. On the one hand, as
452 expected, our results could not identify a factor with high loadings from both self and
453 interpersonal functioning. Some authors argue that interpersonal functioning is not a
454 real independent component, yet is driven by mature self functioning, given having
455 healthy relationships with others requires a well-developed self (Lakuta, 2022). Present
456 results do not support this assumption and overall show distinctiveness in the
457 functioning of both elements of personality impairment. Similarly congruent with our
458 hypothesis, self and interpersonal functioning factors appeared intermingled with, yet
459 also clearly differentially related to, maladaptive domains of AMPD, as operationalized
460 with the PID-SF. Consistent with previous research (Few et al., 2013; Lakuta, 2022;
461 Sleep et al., 2019; Sleep et al., 2020), self functioning was mostly linked to trait of
462 negative affect, with also showing association with psychoticism. interpersonal
463 functioning was grouped with the detachment domain of PID-5.

464 However, some discrepancies with our hypotheses were also found. Our results
465 showed self functioning appears also related with disinhibition. We did not hypothesize
466 this finding, yet Few et al. (2013) did find a medium-sized association of self-
467 directedness, a facet of self functioning and disinhibition. Also, previous studies have
468 found a medium-sized relationship between facets of disinhibition and personality
469 functioning (Clark & Ro, 2014). Further, the antagonistic domain appeared separately in
470 the third factor instead of linked to interpersonal functioning. Similarly, Zimmerman et
471 al. (2015) found that certain facets (manipulativeness, attention seeking and
472 deceitfulness) of antagonism emerged as a separate trait factor. Overall, the results of

473 the joint EFA corroborated previous literature by providing evidence that the two
474 subdimensions of personality functioning combine differentially with the five
475 maladaptive personality domains. Especially the relationship of self functioning with
476 negative affectivity and interpersonal functioning with detachment appear robust. Yet,
477 at the same time some of the medium sized relations observed need extended evidence
478 to further clarify the mixed results over studies.

479

480 *Predictive and incremental validity of LPFS-BF 2.0 (Criterion A) and PID-5-SF*
481 *(Criterion B) scales to explain functional impairment as measured with the*
482 *WHODAS 2.0*

483 Results showed that level personality functioning is related to the degree of
484 functional impairment. Groups formed by differences in LPFS-BF 2.0 total scores
485 indicated that less personality functioning was significantly associated with more
486 functional impairment measured by WHODAS 2.0 scores.

487 Multiple regression analyses also showed a significant contribution of each of
488 two LPFS-BF 2.0 subscales in predicting functional impairment measured by the
489 WHODAS 2.0 (R^2 adjusted = .40). Yet functional impairment was also predicted by
490 traits measured by the PID-5-SF (R^2 adjusted = .45). LPFS-BF 2.0 with only 12 items
491 was thus able to explain a high percentage of variance in functional impairment by
492 itself, compared to the percentage of variance explained by PID-5-BF traits with 100
493 items (R^2 adjusted respectively .40 and .45). We can therefore conclude that although
494 personality functioning only added limited incremental validity (i.e. 2% of variance)
495 above maladaptive personality traits (whereas vice versa 7% of variance was added
496 personality traits above personality functioning), the results indicate a higher efficiency
497 of LPFS-BF 2.0 to explain functional impairment. Previous research yielded mixed

498 results, varying from low incremental validity using traditional PD constructs as
499 outcomes (Bliton et al., 2022; Roche & Jaweed, 2021) but greater contribution when
500 capturing psychological health (Bach & Hutsebaut, 2018). The current results showed
501 the item coverage of LPFS-BF 2.0 is well designed to evaluate patient's overall health
502 and impairment. The limited number of items make the LPFS-BF 2.0 a more effective
503 measure than the PID-5-SF for those contexts where identification of impairment is
504 relevant (and not the type of personality disorder a patient has).

505 It is also important to emphasize that present research suggested that once the
506 personality domains have been introduced into the model, only self functioning remains
507 as an explanatory variable. Congruently, a recently study by Lakuta (2022) retained
508 only the self functioning scale of LPFS-BF 2.0 above the personality domains for
509 explaining well-being in a non-clinical sample. These results suggest that self
510 functioning requires special attention when explaining impairment and psychosocial
511 functioning. Similarly, several studies found the identity component of self functioning,
512 to be able to predict global functioning (Buer Christensen et al., 2020; Esguevillas et
513 al., 2018), and well-being in an older personality disorder patients sample (Veenstra et
514 al., 2022).

515

516 *Clinical implications*

517 From a clinical point of view our work has a number of implications. On one
518 hand the incremental validity of criterion A could be judged to be limited above
519 criterium B, and some therefore even consider measuring personality functioning as
520 redundant (Sleep et al., 2019). Our study results indeed indicate limited incremental
521 validity above traits in predicting functional impairment (and vice versa traits have
522 limited incremental value above personality functioning). On the other hand, both

523 personality functioning and trait domains have substantial incremental validity (40-
524 45%) above gender when predicting functional impairment. Presence of the personality
525 functioning may vary depending on the elevated traits a patient presents (Clark et al.,
526 2020) and, according to some researchers, criterion A can explain the expression of
527 maladaptive traits, and can be interpreted as the underlying impairment in functioning.
528 Overlap is consequently logical, although more conceptual clarity is needed (Sharp &
529 Wall, 2021; Zimmermann et al., 2015). As demonstrated by our joint factor analysis,
530 interpersonal dysfunction is difficult to differentiate from maladaptive traits. Still, our
531 results at the same time indicated self and interpersonal functioning measured with the
532 LPFS-BF 2.0 can be differentiated as demonstrated by the two-factor structure of the
533 LPFS-BF 2.0, and corroborated by their differential loadings and correlations with traits
534 when jointly factor analyzed with maladaptive traits, and by their differential ability to
535 predict impairment. These results lead to consider the LPFS-BF 2.0 a useful tool for
536 clinical routine monitoring. It is a brief measure that can be quickly applied to provide
537 and index of self-reported overall severity of PD. At the same time self and
538 interpersonal functioning can be differentiated and the instrument has adequate
539 explanatory capacity in terms of impairment. The measure has also recently been
540 recommended as part of a standard set of patient-reported outcomes for PD (Prevolnik
541 Rupel et al., 2021).

542

543 ***Limitations***

544 Some limitations of the present study results should also be highlighted. First, it
545 should be noted that in our mixed sample, clinical patients most frequent diagnoses
546 were depressive, anxiety and trauma related disorders. Combined with community data
547 this might imply a lower representation of patients with PD compared to other studies

548 conducted specifically with PD patients, like for example the original validation study
549 of the LPFS-BF 2.0 (Weekers et al., 2019). Although this aspect may limit the
550 representation of the most severe form of traits and dysfunction, high scores of
551 personality impairment in general clinical patients and even in community samples have
552 been reported (Sleep, 2022). In our sample the number of persons with clinical
553 impairment was limited, yet we had scores varying from no impairment, to subclinical,
554 and up to clinical impairment.

555 On the other hand, as reported by Zimmermann et al. (2022), self-report
556 measures likely catch impairment related to mental disorders in general. This relates to
557 the second limitation. Several authors have pointed out that the exclusive use of self-
558 report measures may constitute a limitation in the accuracy of the assessment of
559 personality functioning (Ganellen, 2007; Miller & Lynam, 2015). Although
560 complementing these measures with informant reports could extend the results, clinical
561 practice mostly uses self-report measures, especially for routine patient monitoring.

562 Finally, it should be noted that the sample size of the present work in view of the
563 power analysis performed, did not allow the application of an SEM approach for
564 analysing the incremental validity of the LPFS-BF 2.0 above the PID domains, when
565 predicting functional impairment as measured by the WHODAS 2.0. According to
566 Wang & Eastwick (2020), this may lead to an increase in the Type I error rate, therefore
567 we applied a Bonferonni correction. Additionally, the use of observable scores, instead
568 of latent variables, can be more useful for a direct interpretation of these scores in
569 clinical practice.

570

571 ***Conclusion***

572 Despite the mentioned limitations, we conclude overall results are promising.
573 The LPFS-BF 2.0 can capture general severity and differentiate self and interpersonal
574 functioning. Although personality functioning was defined as a unidimensional severity
575 criterion, self and interpersonal functioning behave as separate and distinct dimensions of
576 personality functioning when measured by LPFS-BF 2.0, and level of personality
577 functioning as measured by this instrument is associated with degree of functional
578 impairment.

579

580 **Acknowledgments**

581 Not applicable

582

583 **Declaration of interest**

584 The authors declare that there are no conflicts of interest to report.

585

586 **Data availability**

587 The data that support the findings of this study are openly available in Arias Montano
588 repository at <http://hdl.handle.net/10272/21157>.

589

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917

919 *Descriptive statistics and independent t-test between subsamples*

	Total sample <i>N</i> = 1074	Community sample <i>n</i> = 717	Clinical sample <i>n</i> = 357			
	<i>M</i> (<i>SD</i>)	<i>M</i> (<i>SD</i>)	<i>M</i> (<i>SD</i>)	<i>t</i>	<i>p</i>	<i>d</i>
LPFS-Self	18.81 (6.41)	8.60 (2.99)	13.18 (4.84)	15.95	<.001	1.08
LPFS-Inter	10.02 (4.23)	8.38 (2.62)	9.69 (3.36)	6.25	<.001	0.44
LPFS total	18.81 (6.41)	16.99 (5.07)	22.87 (7.19)	13.48	<.001	0.91
PID-5 NegAff	8.63 (2.67)	7.77 (2.12)	10.49 (2.69)	16.92	<.001	1.03
PID-5 Detach	6.62 (2.51)	5.89 (1.86)	8.10 (2.95)	12.89	<.001	0.88
PID-5Antag	5.35 (1.65)	5.42 (1.66)	5.20 (1.61)	2.12	.034	0.13
PID-5 Disinh	6.73 (2.31)	6.01 (1.77)	8.16 (2.59)	14.08	<.001	0.92
PID-5 Psycho	6.07 (2.33)	5.40 (1.66)	7.43 (2.85)	12.35	<.001	0.87
WHODAS	20.03 (9.48)	16.33 (6.46)	27.50 (10.19)	18.84	<.001	1.17

920 Anh = Anhedonia; Antag= Antagonism; Anx = Anxiousness; Att = Attention seeking; Call =
921 Callousness; Dec = Deceitfulness; Dep = Depressivity; Detach = Detachment; Dis = Distractibility;
922 Disinh = Disinhibition; Ecc = Eccentricity; Emo = Emotional lability; Gran = Grandiosity; Hos =
923 Hostility; Imp = Impulsivity; Int = Intimacy avoidance; Irr = Irresponsibility; LPFS-Self = Level of
924 Personality Functioning Scale Self Functioning Scale; LPFS-Inter = Level of Personality Functioning
925 Scale Interpersonal Functioning Scale; Man = Manipulativeness; NegAff = Negative Affect; Perc =
926 Perceptual dysregulation; Pers = Perseveration; Psycho = Psychoticism; Res = Restricted affectivity; Rig
927 = Rigid perfectionism; Risk = Risk taking; Sep = Separation insecurity; Sub = Submissiveness; Unu =
928 Unusual beliefs; With = Whitdrawal.

929 Table 2.

930 *Fit statistics for confirmatory factor analyses*

	χ^2 (d.f.)	CFI	TLI	SRMR	RMSEA [90% CI]
<i>Model 1: 1 factor</i>	698.217 (54)	.948	.937	.074	.105 [.098; .112]
<i>Model 2: 2 correlated factors</i>	338.18 (53)	.977	.972	.049	.071 [.064; .078]
<i>Model 3: Bi-factor</i>	132.742 (42)	.993	.989	.028	.045 [.036; .054]

931 *Note.* Estimation Method: Diagonally Weighted Least Squares Mean Variance Adjusted (WLSMV); CFI =
932 Comparative Fit Index; TLI = Tucker Lewis Index; RMSEA = Root Mean Square Error Approximation; SRMR
933 = Standardized Root Mean Square Residual. All chi-square values are statistically significant.

Table 3. *Fit indices for different factor solutions of exploratory factor analysis*

	χ^2 (d.f.)	CFI	TLI	SRMR	RMSEA [90% CI]
1 factor	329.432 (14)	0.921	0.882	0.052	0.145 [0.132 – 0.159]
2 factors	122.646	0.971	0.925	0.029	0.116 [0.098 – 0.134]
3 factors	6.477 (3)	0.998	0.994	0.005	0.033 [0.001 – 0.068]

Table 4. *Factor loadings the Joint Exploratory Factor Analysis of the LPFS-BF 2.0 (Criterion A) and PID-5-SF (Criterion B) (N = 1074)*

	Factor 1	Factor 2	Factor 3
NA	.891	-.119	-.006
DET	.010	.882	.001
ANT	-.017	-.005	.742
DIS	.617	.089	.259
PSY	.438	.295	.278
LPFS-Self	.704	.296	-.117
LPFS-Inter	.188	.422	.289

Note. Factor loadings above .30 are indicated in bold
 ANT = Antagonism; DET = Detachment; DIS = Disinhibition; NA = Negative Affect; PSY = Psychoticism; LPFS-Self = LPFS-BF 2.0 Self functioning; LPFS-Inter = LPFS-BF 2.0 Interpersonal functioning

Table 5.

Model A: Multiple regression analysis predicting WHODAS 2.0 scores (Criterion A second step).

		Coef.	Std. Err.	Beta	t	P > t	I.C. 95%		Adjusted R ²
								Lím. Inf.	Lím. Sup.
Step 1	Gender (female)	3.23	0.58	.17	5.49	< .001	2.07	4.38	.03
	Age	-0.03	0.02	-.04	-1.43	.151	-0.07	0.01	
Step 2	Gender (female)	1.51	0.48	.08	3.18	.002	.58	2.45	.40
	Age	0.07	0.02	.10	4.12	< .001	.03	.10	
	Self functioning LPFS-BF 2.0	1.30	0.07	.57	18.64	< .001	1.16	1.43	
	Interpersonal functioning LPFS-BF 2.0	0.36	0.11	.10	3.42	.001	.15	.57	
Step 3	Gender (female)	0.64	0.46	.03	1.40	.163	-.26	1.55	.47
	Age	0.04	0.02	.07	2.90	.004	.01	.07	
	Self functioning LPFS-BF 2.0	0.59	0.09	.26	6.41	< .001	.41	.77	
	Interpersonal functioning LPFS-BF 2.0	0.01	0.11	.00	0.10	.922	-.21	.24	
	Negative affect	0.63	0.12	.18	5.35	< .001	.40	.86	
	Detachment	0.92	0.13	.24	7.31	< .001	.67	1.17	

Antagonism	-0.83	0.15	-.15	-5.54	< .001	-1.13	-.54
Disinhibition	0.37	0.15	.09	2.52	.012	.08	.66
Psychoticism	0.34	0.15	.08	2.26	.024	.04	.63

$R^2 = .47$; R^2 adjusted = .47; $F(9, 1056) = 105.01$; $p < .001$; $LL = -3567.84$; $LR = 132.04$; $p < .001$; $N = 1074$

Table 6.

Model B: Multiple regression analysis predicting WHODAS 2.0 scores (Criterion B second step).

		Coef.	Std. Err.	Beta	t	P > t	I.C. 95%		Adjusted R ²
								Lím. Inf.	Lím. Sup.
Step 1	Gender (female)	3.23	0.58	.17	5.49	< .001	2.07	4.38	.03
	Age	-0.03	0.02	-.04	-1.43	.151	-0.07	0.01	
Step 2	Gender (female)	0.78	0.46	.04	1.69	.092	-.13	1.69	.45
	Age	0.03	0.02	.04	1.64	.102	-.01	.06	
	Negative affect	0.92	0.11	.26	8.30	< .001	.70	1.13	
	Detachment	1.23	0.11	.33	10.86	< .001	1.01	1.46	
	Antagonism	-0.99	0.15	-.17	-6.67	< .001	-1.28	-.70	
	Disinhibition	0.57	0.15	.14	3.88	< .001	.28	.85	
	Psychoticism	0.53	0.15	.13	3.60	< .001	.24	.82	
	Gender (female)	0.64	0.46	.03	1.40	.163	-.26	1.55	.47
	Age	0.04	0.02	.07	2.90	.004	.01	.07	
	Self functioning LPFS-BF 2.0	0.59	0.09	.26	6.41	< .001	.41	.77	

Step 3	Interpersonal functioning LPFS-BF 2.0	0.01	0.11	.00	0.10	.922	-.21	.24
	Negative affect	0.63	0.12	.18	5.35	< .001	.40	.86
	Detachment	0.92	0.13	.24	7.31	< .001	.67	1.17
	Antagonism	-0.83	0.15	-.15	-5.54	< .001	-1.13	-.54
	Disinhibition	0.37	0.15	.09	2.52	.012	.08	.66
	Psychoticism	0.34	0.15	.08	2.26	.024	.04	.63

$R^2 = .47$; R^2 adjusted = .47; $F(9, 1056) = 105.01$; $p < .001$; $LL = -3567.84$; $LR = 132.04$; $p < .001$; $N = 1074$

Figure 1.

Confirmatory factor analysis: The two-correlated factor model with completely standardized loadings

